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**Supply Capacity, Vertical Specialisation and Trade Costs:  
The Implications for Aggregate US Trade Flow Equations**

Menzie D Chinn (University of Wisconsin, Madison and NBER)

**WORKING PAPER SERIES**

Centre for Competitive Advantage in the Global Economy

Department of Economics

# **Supply Capacity, Vertical Specialization and Trade Costs: The Implications for Aggregate U.S. Trade Flow Equations**

by

Menzie D. Chinn  
University of Wisconsin, Madison  
and NBER

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## **Abstract**

This paper re-examines aggregate and disaggregate import and export demand functions for the United States over the 1975q1-2010q1 period. This re-examination is warranted because (1) income elasticities are too high to be warranted by standard theories, and (2) remain high even when it is assumed that supply factors are important. These findings suggest that the standard models omit important factors. An empirical investigation indicates that the rising importance of vertical specialization combined with changing tariff rates and transportation costs explains some of results. Accounting for these factors yields more plausible estimates of income elasticities.

**Keywords:** imports, exports, elasticities, vertical specialization, production fragmentation, trade costs

**JEL Classification:** F12, F41

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**Correspondence:** Robert M. LaFollette School of Public Affairs; and Department of Economics, University of Wisconsin, 1180 Observatory Drive, Madison, WI 53706-1393. Tel.: +1 (608) 262-7397. Email: [mchinn@lafollette.wisc.edu](mailto:mchinn@lafollette.wisc.edu) .

## 1. Introduction

This analysis is motivated by the recent widening and narrowing of the U.S. trade balance, illustrated in Figure 1, and three key stylized facts regarding U.S. trade flows. The first stylized fact is the persistence of the Houthakker-Magee results, namely that the income elasticity of U.S. imports exceeds that of exports. Over the past 30 years, the gap is at least 0.3 for total goods and services, regardless of the method of estimation. In Chinn (2005), the gap is as high as 0.65. Table 1 presents estimates obtained from OLS, dynamic OLS, single equation error correction estimates and the Johansen maximum likelihood procedures confirming that finding. Furthermore, there is little evidence that the asymmetry is disappearing.

The second stylized fact is that values of the income elasticities are quite high, and rising. In Table 1, the income elasticities are as high as 2.2 for imports, and 1.9 for exports. In Figure 2, the income elasticities are shown for four subperiods over the last 32 years into four equal sub-periods. The income elasticities are generally rising. Notice that in Figure 3 there is no similar pattern for price elasticities. One outstanding feature of the estimates is that the import price elasticity is typically quite low.

These high income elasticities are difficult to reconcile with the standard differentiated goods model (see Goldstein and Khan, 1985). From a forecasting standpoint, high income elasticities<sup>1</sup> are not troubling; but – as discussed below – from an economic perspective, they are perplexing. The third stylized fact is that the variability of trade flows is increasing over time. For instance, the behavior of trade flows during 1999-

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<sup>1</sup> This phenomenon has been noted before (Rose, 1991).

2000 is difficult to explain using standard models. As illustrated in Figure 3, both series surge in this period.

In this paper, I re-examine the behavior of export and import flows, motivating the analysis by referring to new theories of trade behavior. These include differentiated goods models such as those forwarded by Krugman (1991). Several papers have exploited the implications of these models in a cross-section context, with some success. In this paper, I adopt a different approach. Disaggregating the data, one finds that some of the odd behavior of goods exports and imports can be isolated to the peculiar behavior of capital and durable goods; since such goods are often used to manufacture other capital goods or consumer goods, it seems that growth in such categories is “inflating” the volume of such trade flows. Disaggregation is not sufficient, however. Various papers have pointed out that the growth of trade in intermediate goods may be nonlinearly related to the decline in trade barriers and the heightened importance of capital expenditures during certain phases of the business cycle. More recently, Mann and Plück (2007) have argued that disaggregation along category line and trading partner helps in obtaining reasonable parameter estimates.

Once one includes the variables that should matter for such vertical specialization, the parameter estimates become more plausible. That being said, the parameter estimates for the auxiliary variables are not always in the expected direction or statistically significant, and the results cannot be construed as definitive.

## 2. The Standard Model and the Supply Side

### 2.1 The model specification

The empirical specification is motivated by the traditional, partial equilibrium view of trade flows. Goldstein and Khan (1985) provide a clear exposition of this “imperfect substitutes” model. To set ideas consider the algebraic framework similar to that used by Rose (1991). Demand for imports in the US and the Rest-of-the-World (RoW) is given by:

$$D_{im}^{US} = f_1^{US}(Y^{US}, \hat{P}_{im}^{US}) \quad (1)$$

$$D_{im}^{RoW} = f_1^{RoW}(Y^{RoW}, \hat{P}_{im}^{RoW}) \quad (2)$$

where  $\hat{P}_{im}$  is the price of imports relative to the economy-wide price level. The supply of exports is given by:

$$S_{ex}^{US} = f_2^{US}(\hat{P}_{ex}^{US}, Z^{US}) \quad (3)$$

$$S_{ex}^{RoW} = f_2^{RoW}(\hat{P}_{ex}^{RoW}, Z^{RoW}) \quad (4)$$

Where  $\hat{P}_{ex}$  is the price of exports relative to the economy-wide price level. Note that the price of imports into the US is equal to the price of foreign exports adjusted by the real exchange rate.

$$\hat{P}_{im}^{US} \times P^{US} = E \times \hat{P}_{ex}^{RoW} \times P^{RoW} \Rightarrow \hat{P}_{im}^{US} = Q \hat{P}_{ex}^{RoW} \quad (5)$$

where  $E$  is the nominal exchange rate in US\$ per unit of foreign currency, the real exchange rate is

$$Q = \frac{EP^{RoW}}{P^{US}}$$

where  $P$  represents the *aggregate* level of prices of domestically produced goods and services.  $Z$  is a supply shift variable, representing the productive capacity of the exportables sector.

An analogous equation applies for imports into the rest-of-the-world. Imposing the equilibrium conditions that supply equals demand, one can write out import and export equations (assuming log-linear functional forms, where lowercase letters denote log values of upper case):<sup>2</sup>

$$im_t = \beta_0 + \beta_1 q_t + \beta_2 y_t^{US} + \beta_3 z^{RoW} + \varepsilon_{2t} \quad (6)$$

$$ex_t = \delta_0 + \delta_1 q_t + \delta_2 y_t^{RoW} + \delta_3 z^{US} + \varepsilon_{1t} \quad (7)$$

Where  $\beta_1 < 0$ ,  $\beta_2 > 0$ ,  $\beta_3 > 0$  and  $\delta_1 > 0$ ,  $\delta_2 > 0$ ,  $\delta_3 > 0$ .

Notice that exports are the residual of production over domestic consumption of exportables; similarly imports are the residual of foreign production over foreign consumption of tradables. The difference between this specification and the standard is the inclusion of the exportables supply shift variable,  $z$ . In standard import and export regressions, this term is omitted, implicitly holding the export supply curve fixed; in other words, it constrains the relationship between domestic consumption of exportables and production of exportables to be constant (see Helkie and Hooper, 1988 for an exception to this rule). A bout of consumption at home that reduces the supply available for exports would induce an apparent structural break in the equation (6) if the  $z$  term is

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<sup>2</sup> As Marquez (1994) has pointed out, there are a number of problems with this specification, in terms of assumptions regarding expenditure shares. A number of other potentially important factors are also omitted, including other trend variables (e.g, immigration as in Marquez (2002) or the rise of services exports as in Mann (1999)).

omitted. Similarly, omission of the rest-of-world export supply term from the import equation makes the estimated relationships susceptible to structural breaks.

Note that the supply term here is explicitly partial equilibrium in nature. Unlike the Krugman (1991) model, where balanced trade implies supply creates its own demand, no specific presumptions are made regarding the source of this supply effect.

The problem, of course, is obtaining good proxies for these supply terms. In the some previous studies, a measure of the U.S. capital stock has been used. Obvious candidates, such as US industrial production for US exports, exhibits too much collinearity with rest-of-world GDP to identify the supply effect precisely. That is why this supply factor has typically been identified in panel cross section analyses (Bayoumi, 2003; Gagnon, 2004).

## ***2.2 Data and Estimation***

Data on real imports and exports and components of real GDP (2005 chain weighted dollars) were obtained for the 1967q1-2010q1 period. Domestic economic activity is measured by U.S. GDP in 2005 chain weighted dollars. Foreign economic activity is measured by real Rest-of-World GDP, weighted by U.S. exports to major trading partners. The real exchange rate measure is the Federal Reserve Board's broad trade weighted value of the dollar. This index uses the CPI as the deflator.



The supply variables are more difficult to identify. For the United States, two candidate variables are used: (i) net private nonresidential capital, and (ii) manufacturing production.<sup>3</sup> For the rest-of-the-world, I use import-weighted rest-of-world GDP.

The trade costs are proxied by two variables: (i) the average tariff rate for major economies, and (ii) the relative price of oil. (Additional details on all these variables are contained in Appendix 1.)

Estimation is implemented on data over the period of 1975q1-2010q1. This period spans three episodes of dollar appreciation and three episodes of dollar depreciation. It also spans a period of tremendous volatility in trade flows, from 2008q3-09q3; in order to maximize the span of data, I retain this data for the analysis. In addition, the variation in the series should aid in identifying the relationships (as long as there are no structural breaks associated with this the trade collapse).

The broad measure of dollar is used, as opposed to the major currencies measure, which is unrepresentative of relative prices faced by the U.S. import competing sector in recent years. In principle, it would be better to use a unit labor deflated index (Chinn, 2006). However, such an index is not available for a broad set of currencies; of which the Chinese yuan would be the most important.<sup>4</sup> In addition, by using this Divisia exchange rate index (which weights *growth rates* of exchange rates by trade weights), I sidestep the

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<sup>3</sup> Admittedly, industrial production is in some sense too “endogenous” a variable to include in the regression. From a theoretical perspective, a more desirable measure is the capital stock measures as a proxy of the supply capacity of exportables, as in Helkie and Hooper (1988). The question is whether these variables are measured with too much error. Certainly, this is an almost impossible task for the rest-of-the-world, especially to the extent that one wants to capture the impact of the newly industrializing countries and China.

<sup>4</sup> For an analysis of how different results are obtained using different deflators, see Chinn (2005). In that analysis, I examine the impact of using alternatively CPI, PPI and unit labor cost deflated real exchange rates on estimates of trade elasticities.

question of whether the impact of Chinese relative prices are fully captured. For more this issue, see Thomas and Marquez (2006).

In terms of estimation, there are a variety of alternative approaches to estimating the long run relationships, include the Johansen maximum likelihood method, and single equation error correction regressions. The advantage of such approaches is that they provide information on short run dynamics. One of the disadvantages is that the estimates are extremely sensitive to the treatment of trends, and the selection of lag length.<sup>5</sup>

As a consequence, the cointegrating relationship is identified using dynamic OLS (Stock and Watson, 1993). Two leads and four lags of the right hand side variables are included. In a simple two variable cointegrating relationship, the estimated regression equation is:

$$y_t = \gamma_0 + \gamma_1 x_t + \sum_{i=2}^{-4} \Gamma_i \Delta x_{t+i} + u_t$$

This approach presupposes that there is only one long run relationship. This requirement should not be problematic, as there is typically little evidence of more than one cointegrating vector for these four-variable systems.

## ***2.3 Empirical Results***

### ***2.3.1 The Basic Specification***

First I consider equations (6) and (7) suppressing the  $z$  terms. The long run export elasticities are reported in Table 2. The income elasticity for total exports of goods and services is 1.81 (Column [1]). This finding is not an artifact of the inclusion of services. In fact, the goods only elasticity is 1.87 (Column [3]). The high coefficients hold for

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<sup>5</sup> See Chinn (2005) for results using these approaches.

different aggregates including goods ex Agricultural goods. The services elasticities are slightly lower, at about 1.67.

A similar result obtains for imports. As reported in Table 3, Column [1], total imports of goods and services also exhibit a long run elasticity of 2.19. However, in all instances, there appears to be substantial sensitivity to the inclusion of time trends, suggesting the omission of important variables. (Alternatively, a cointegrating relationship is absent between these specific trade aggregates and the included income and price variables.)

In addition to the empirical motivation for examining different aggregates, there is a standard convention to consider, for instance, an import aggregate excluding petroleum. The trade equations in (6) and (7) are derived from an imperfect substitutes model, well suited to manufactured goods. However, oil is a natural resource commodity that does not quickly respond to market signals, and exhibits trends due to resource depletion. A similar argument might be used to motivate a focus on a non-agricultural goods export variable. Figures 5 and 6 depict these alternative series and break out services.

Interestingly, these remarkably high estimated income elasticity estimates persist, for both export goods ex.-Agriculture and import goods ex.-oil (columns [3] in Tables 2 and 3). On the other hand, estimated price elasticities *are* higher (in absolute value) for these disaggregate components. These findings suggest some aggregation bias.

These high estimated income elasticities inform the debate over the durability of the Houthakker-Magee (1969) findings. Exports involving goods respond 1.8 to 1.9 percent for each one percentage point increase in rest-of-world income. In contrast, imports rise about 2.3 to 2.6 percentage points for each percentage point increase in US

GDP. This set of findings suggests that the Houthakker-Magee income asymmetry persists. Hence, even if U.S. and foreign growth rates were to converge, net exports would continue to deteriorate even starting from balanced trade.

### 2.3.2 A Supply Augmented Specification

All of the preceding specifications exclude a role for the supply side, suggested by Equations (6) and (7). In Table 4, the results for specifications incorporating the supply side are included, where the supply side is proxied by manufacturing production.<sup>6</sup> The income elasticity of demand falls from 1.8 to 0.9, with the supply coefficient equal to close to unity (Column [1]). Unfortunately, the results are sensitive to the inclusion of a time trend. Counter-intuitively, exports of goods ex Agriculture are not as easily modeled for this specification; the rest-of-world income is not statistically significant. Inclusion of a time trend leads to a negative coefficient estimate on real GDP.

On the import side, the inclusion of supply side effects is slightly less successful. In this case, the supply side variable is import-weighted real GDP. In Table 4, Column [5], the import income elasticity rises from 2.19 to 4.07, with the coefficient on foreign GDP taking on a negative value. Excluding oil from the import aggregate leads to slightly more sensible results, but the supply side variable is never significant.

Unfortunately, in all these instances, the demand and supply variables are so collinear that the results are sensitive to the inclusion of the time trends.<sup>7</sup> This is why

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<sup>6</sup> The regressions incorporating private nonresidential capital stock are unsuccessful to the extent that the coefficient on the supply variable exhibits the wrong sign. Interestingly, the two series exhibit similar trends, and according to certain tests, are cointegrated.

<sup>7</sup> For a survey of research related to the ongoing trade collapse and rebound, see Baldwin (2009). Freund (2009) provides empirical evidence over time.

cross-section and panel regressions such as Gagnon (2003) and Bayoumi (2003) obtain more supportive evidence of supply side effects.

### **3. Vertical Specialization and Tariffs**

One hint of why the income elasticities are so large is provided by the surge in both exports and imports during 1999-2000. In informal discussions, this jump is associated with the investment boom; the category experiencing the largest jump is capital goods. The fact that the surge and collapse occurred in both categories could be coincidence – evidence of a synchronized worldwide investment boom. Or it could be a reflection that the two are interlinked.

Recent research has focused on the rise of intermediate goods in international trade. However, intermediate goods are not in and of themselves sufficient to explain the rise in trade. It is intermediate goods trade used to produce other traded goods – in other words vertical specialization (Hummels, et al., 2001; Yi, 2003; Chen et al., 2005) – that is required. This process of importing in order to export has also been termed the “fragmentation” of the production process (Arndt, 1997). At this juncture, it is useful to recognize that services exhibit less of this fragmentation.<sup>8</sup> This explains in part the differential import income elasticities: 2.62 for goods ex oil versus 1.64 for services.<sup>9</sup>

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<sup>8</sup> Barrell and Dées (2005) and Camerero and Tamarit (2003) address the issue of very high income elasticities by incorporating FDI into the specifications. IMF (2007) incorporates exports of intermediates in the import equation, and imports of intermediates in the export equation, to account for vertical specialization. This procedure reduces the estimated income elasticities.

<sup>9</sup> Marquez (2005) obtains similar estimates, but points out that further disaggregation of services leads to different insights on income and price elasticities.

Figures 7 and 8 show how differently different goods aggregates behave. Note that the series excluding capital goods exhibits much less of a pronounced hump. The importance of vertical specialization was suggested, particularly for hi-tech goods, in analyses around the time of the capital goods surges (e.g., Council of Economic Advisers, 2001, Chapter 4).

The regression results in Column [1] of Table 5 repeat the results in Column [3] of Table 4. Adding in the trade costs – tariffs and transportation costs – leads to substantially improved results. Foreign demand and home supply enter with roughly equal coefficients, while the price elasticity is fairly high, at 0.744. The tariff factor and the square both enter with statistical significance, indicating that lower tariffs increase trade flows. However, as expected higher energy costs, as proxied by the relative price of oil, also enters in. This result corroborates the findings of Bergin and Glick (2006). Qualitatively, the results are the same when examining the goods ex.-agriculture, ex.-capital goods (column [3]), with the exception of a role for transport costs.

An alternative breakdown is between durables and nondurables. Durables exports and imports are graphed in Figure 9, along with the tariff factor. Notice that as the tariff factor flattens out, the trend growth rate of durables growth slows down. The estimates for durables and nondurables aggregates on the export side are reported in columns [4] and [5], respectively. Durables exhibit the posited relationship, while nondurables are not apparently related to the supply variable, and transport costs.<sup>10</sup> Finally, in line with expectations, the coefficient on foreign income is the highest for the capital goods exports

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<sup>10</sup> In addition, the tariff factor coefficients in the full specification exhibit the wrong signs. I drop the level variable in the reported results.

category. In this case, a one percent increase in foreign income is associated with a 1.3 percent increase in capital goods exports.

In general, it is more difficult to fit the import data. The results in Column [8] contrasts with the baseline results in Column [7], insofar as the income elasticities (for demand and supply) are much more plausible. The coefficients on tariffs and transportation costs enter with the correct sign, although the latter coefficient is not statistically significant. Qualitatively similar results obtain for imports of goods ex.-oil, ex.-capital goods.

Using the alternative breakdown, one finds that the coefficients for the durables aggregate are correctly signed, although transportation costs are not significant. For nondurables, foreign supply, the exchange rate and transportation costs fail to show up as statistically significant.

Finally, capital goods imports are particularly difficult to model. Most of the variables are not statistically significant or implausibly large in absolute value. The standard error of regression is also large (0.061) relative to that for the other regressions for the other aggregates. A similar result obtains for capital goods imports, where the SER is 0.055.

One feature of the results in Table 5 is that when disaggregating, allowing for supply effects, and changing trade costs, then the estimated price elasticities are typically higher than for the specifications reported in Tables 2 and 3.<sup>11</sup> Future work will involve aggregating up the price elasticities for the components to obtain a more accurate measure of the aggregate price elasticity.

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<sup>11</sup> One exception is for capital goods imports. In that case, a statistically significant price elasticity cannot be obtained.

#### **4. Summary**

In this paper the data for U.S. trade flows up to 2010q1 are investigated. A variety of different trade aggregates are examined. In addition, supply side factors and the implications of vertical specialization are accounted for. A number of conclusions are derived from this assessment.

First, the examination confirms that the Houthakker-Magee finding of income asymmetries persists into the most recent period. This characterization applies most strongly to specifications involving highly aggregated trade flows, and no role for supply or other factors.

Second, the disaggregation of trade flows into services and some subcategory of goods usually yields higher estimated price elasticities. This outcome suggests some role for aggregation bias in driving down estimated price elasticities. A similar finding was obtained in IMF (2007), but in that case, the results pertained to price elasticities for relative prices, instead of the real exchange rate as in this study.

Third, the inclusion of supply-side variables reduces the magnitude of income elasticities for goods. However, the results are not robust to the inclusion of time trends. Consequently, one can only make tentative conclusions regarding the importance of supply side factors in driving the increasing volume of international trade. On the other hand, cross-section studies of trade do suggest that the finding of a supply side role is not completely coincidental.

Fourth, capital goods and non-capital goods imports, or alternatively, durable and nondurable goods, appear to behave differently. However, because the results are sensitive to the sample period and trade flow measure, additional work is required to



identify the channels by which trade barriers and vertical specialization interact. In particular, one might want a better measure of trade barriers for trade in capital goods.

Finally, it appears that disaggregation – even of a limited extent – might prove helpful in improving predictions of aggregate trade flows. One key dividing line appears to be between non-oil non-capital goods and capital goods, and between durables and nondurables (particularly on the export side).<sup>12</sup>

The results so far suggest further avenues of research.

1. Checking the sensitivity of the results to alternative specifications, including additional activity variables pertaining to specific trade flows (e.g., capital goods imports as a function of business fixed investment).
2. Investigating the results obtained from the Johansen methodology and single equation error correction specifications.
3. Incorporating different or estimated measures of trade costs.
4. Focus on the short run dynamics of the components of trade flows, such as durables versus nondurables. This point is related to item (1), since there is some evidence that the extreme movements in durable goods was associated with a particularly sharp movement in (highly procyclical) durable good production (Francois and Woerz, 2009).

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<sup>12</sup> In previous versions of this paper, I've reported results that indicate that sums of predicted import sub-aggregates appear to yield smaller prediction errors than using a predicted aggregate import variable. This finding – while not definitive – suggests that one can improve our forecasts of trade flows without resorting to modeling many very highly disaggregated trade series.

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**Table 1: Estimates of Export and Import Elasticities, 1975q1-2010q1**

	Exports of Goods and Services				Imports of Goods and Services			
	OLS	DOLS <sup>a/</sup>	ECM <sup>b/</sup>	VECM	OLS	DOLS <sup>a/</sup>	ECM <sup>b/</sup>	VECM
	[1]	[2]	[3]	[4]	[5]	[6]	[7]	[8]
Income (Demand)	<b>1.780</b>	<b>1.811</b>	<b>1.820</b>	<b>1.991</b>	<b>2.174</b>	<b>2.190</b>	<b>2.171</b>	<b>2.222</b>
	<b>[0.032]</b>	<b>[0.031]</b>	<b>[0.056]</b>	<b>[0.042]</b>	<b>[0.036]</b>	<b>[0.028]</b>	<b>[0.052]</b>	<b>[0.035]</b>
Exchange rate	<b>0.435</b>	<b>0.568</b>	<b>0.944</b>	<b>0.947</b>	<b>-0.197</b>	-0.151	<b>-0.308</b>	<b>-0.163</b>
	<b>[0.094]</b>	<b>[0.095]</b>	<b>[0.259]</b>	<b>[0.166]</b>	<b>[0.074]</b>	[0.086]	<b>[0.200]</b>	<b>[0.126]</b>
Adj. R2	0.99	0.99	0.37	Na	0.99	0.99	0.41	Na
SER	0.066	0.052	0.019	Na	0.055	0.047	0.024	Na
N	140	138	141	141	141	139	141	141
Coint. Vectors	Na	na	1	1,1	na	na	1	1,1

**Notes:** Point estimates and HAC standard errors for OLS and DOLS in [brackets], implied long run coefficients from ECM and cointegrating vector coefficients for VECM [asymptotic standard errors in brackets]. SER is standard error of regression. N is number of observations.

Cointegrating vectors is the number of indicated cointegrating vectors; under VECM, {#,#} indicates the number of vectors as indicated by the trace and maximal eigenvalue statistics at the 5% level, using the asymptotic critical values. **[bold face]** indicates significance at the 10% level.

<sup>a/</sup> Includes 2 leads and 4 lags of the first differenced right hand side variables.

<sup>b/</sup> Includes 3 lags of the first differenced variables.

**Table 2: Export Equations, 1975q1-2010q1**

	Total goods & svcs.	Total goods & svcs.	Total goods	Total goods	Total goods ex. Agric.	Total goods ex. Agric.	Total svcs.	Total svcs.
	[1]	[2]	[3]	[4]	[5]	[6]	[7]	[8]
Income (Demand)	<b>1.811</b> [0.032]	<b>3.795</b> [0.783]	<b>1.873</b> [0.037]	<b>4.153</b> [1.051]	<b>1.848</b> [0.054]	<b>3.454</b> [1.782]	<b>1.666</b> [0.030]	<b>2.883</b> [0.682]
Exchange Rate	<b>0.568</b> [0.095]	<b>0.399</b> [0.092]	<b>0.591</b> [0.128]	<b>0.398</b> [0.123]	<b>0.826</b> [0.205]	<b>0.690</b> [0.221]	<b>0.467</b> [0.095]	<b>0.364</b> [0.101]
time		<b>-0.017</b> [0.006]		<b>-0.019</b> [0.009]		-0.013 [0.015]		<b>-0.010</b> [0.006]
Adj. R2	0.99	0.99	0.99	0.99	.98	98	0.99	0.99
SER	0.052	0.048	0.064	0.060	0.094	0.093	0.052	0.050
N	138	138	138	138	138	138	138	138

**Notes:** Point estimates and HAC standard errors for OLS and DOLS in [brackets]. SER is standard error of regression. N is number of observations. Regressions include 2 leads and 4 lags of first differenced right hand side variables. **[bold face]** indicates significance at the 10% level.

**Table 3: Import Equations, 1975q1-2010q1**

	Total goods & svcs.	Total goods & svcs.	Total goods	Total goods	Total goods ex oil	Total goods ex oil	Total svcs.	Total svcs.
	[1]	[2]	[3]	[4]	[5]	[6]	[7]	[8]
Income (Demand)	<b>2.190</b> <b>[0.028]</b>	<b>2.891</b> <b>[0.337]</b>	<b>2.306</b> <b>[0.035]</b>	<b>3.221</b> <b>[0.434]</b>	<b>2.612</b> <b>[0.017]</b>	<b>2.537</b> <b>[0.332]</b>	<b>1.650</b> <b>[0.029]</b>	<b>1.219</b> <b>[0.404]</b>
Exchange Rate	<b>-0.151</b> <b>[0.086]</b>	<b>-0.138</b> <b>[0.067]</b>	-0.120 [0.116]	-0.103 [0.094]	<b>-0.445</b> <b>[0.075]</b>	<b>-0.446</b> <b>[0.075]</b>	<b>-0.289</b> <b>[0.110]</b>	<b>-0.296</b> <b>[0.106]</b>
time		<b>-0.005</b> <b>[0.003]</b>		<b>-0.007</b> <b>[0.003]</b>		0.001 [0.003]		0.003 [0.003]
Adj. R2	0.99	0.99	0.99	0.99	0.99	0.99	0.99	0.99
SER	0.047	0.045	0.061	0.059	0.038	0.038	0.056	0.056
N	139	139	139	139	139	139	139	139

**Notes:** Point estimates and HAC standard errors for OLS and DOLS in [brackets]. SER is standard error of regression. N is number of observations. Regressions include 2 leads and 4 lags of first differenced right hand side variables. **[bold face]** indicates significance at the 10% level.

**Table 4: Supply Augmented Specifications, 1975q1-2010q1**

	Exports				Imports			
	Total goods, supply side	Total goods, supply side	Total goods, ex. Agric., Supply side	Total goods, ex. Agric., Supply side	Total goods supply side	Total goods supply side	Total goods ex Oil, supply side	Total goods ex Oil, supply side
	[1]	[2]	[3]	[4]	[5]	[6]	[7]	[8]
Income (Demand)	<b>0.890</b> [0.184]	<b>2.451</b> [0.787]	0.193 [0.178]	-0.111 [4.012]	<b>4.073</b> [0.308]	<b>3.707</b> [0.267]	<b>2.213</b> [0.312]	<b>2.251</b> [0.323]
Output (Supply)	<b>1.048</b> [0.186]	<b>0.977</b> [0.177]	<b>1.766</b> [0.178]	<b>1.789</b> [0.198]	<b>-1.416</b> [0.261]	<b>-4.711</b> [0.783]	0.328 [0.248]	0.672 [0.598]
Exchange Rate	<b>0.711</b> [0.080]	<b>0.575</b> [0.101]	<b>1.026</b> [0.073]	<b>1.052</b> [0.117]	0.049 [0.077]	<b>0.386</b> [0.094]	<b>-0.477</b> [0.081]	<b>-0.508</b> [0.094]
time		<b>-0.013</b> [0.016]		0.002 [0.007]		<b>0.034</b> [0.008]		0.006 [0.006]
Adj. R2	0.99	0.99	0.99	0.99	.99	.99	0.99	0.99
SER	0.044	0.042	0.048	0.048	0.052	0.039	0.037	0.037
N	138	138	138	138	138	138	138	138

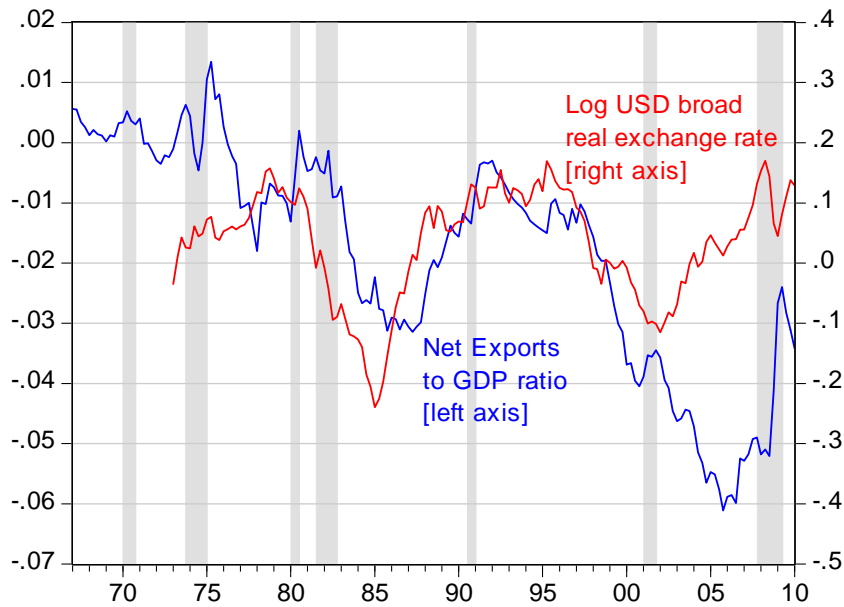
**Notes:** Point estimates and HAC standard errors for OLS and DOLS in [brackets]. SER is standard error of regression. N is number of observations. Regressions include 2 leads and 4 lags of first differenced right hand side variables. **[bold face]** indicates significance at the 10% level.



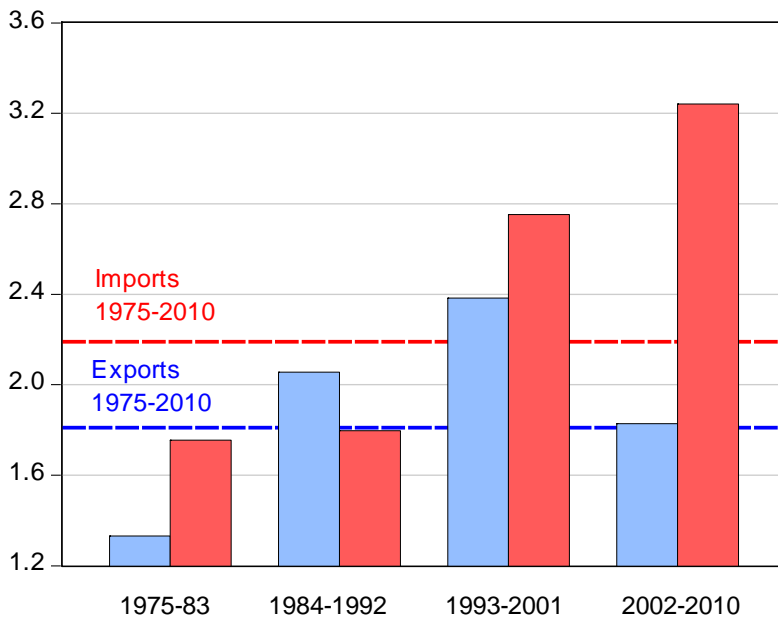
**Table 5: Vertical Specialization and Trade Flows: 1975q1-2010q1**

	Exports of goods ex. Agric.	Exports of goods ex. Agric.	Exports of goods ex Agric., Capital goods	Exports of Durable goods	Exports of Non-durable goods	Exports of Capital goods	Imports of goods ex oil	Imports of goods ex Oil	Imports of goods ex Oil, ex Capital goods	Imports of Durable goods	Imports of Non-durable goods	Imports of Capital goods
	[1]	[2]	[3]	[4]	[5]	[6]	[7]	[8]	[9]	[10]	[11]	[11]
Income (Demand)	0.193 [0.178]	<b>0.920</b> [0.217]	<b>0.671</b> [0.207]	<b>1.293</b> [0.234]	<b>0.787</b> [0.189]	<b>1.341</b> [0.284]	<b>2.213</b> [0.312]	<b>0.903</b> [0.448]	<b>1.736</b> [0.386]	<b>1.047</b> [0.586]	<b>2.836</b> [0.632]	-0.618 [0.880]
Income (Supply)	<b>1.766</b> [0.178]	<b>0.744</b> [0.257]	<b>0.760</b> [0.253]	<b>0.552</b> [0.247]	0.081 [0.111]	<b>0.657</b> [0.314]	0.328 [0.248]	<b>1.033</b> [0.330]	0.206 [0.270]	<b>0.988</b> [0.433]	-0.703 [0.457]	<b>2.842</b> [0.652]
Exchange Rate	<b>1.026</b> [0.073]	<b>0.914</b> [0.150]	<b>0.747</b> [0.120]	<b>0.916</b> [0.187]	<b>0.541</b> [0.126]	<b>1.131</b> [0.228]	<b>-0.477</b> [0.081]	<b>-0.428</b> [0.104]	<b>-0.560</b> [0.088]	<b>-0.365</b> [0.131]	0.040 [0.163]	-0.084 [0.192]
Tariff rate		<b>-389.49</b> [119.88]	<b>-392.14</b> [106.74]	<b>-332.89</b> [128.12]	-- --	<b>-316.81</b> [165.11]		<b>-208.65</b> [86.02]	-88.78 [71.70]	<b>-192.75</b> [111.19]	<b>-525.44</b> [129.59]	240.18 [166.65]
Tariff rate (sq.)		<b>183.34</b> [55.43]	<b>188.71</b> [49.71]	<b>154.80</b> [59.13]	<b>-2.887</b> [1.107]	<b>143.59</b> [76.18]		<b>95.48</b> [39.84]	42.18 [32.91]	-86.55 [51.16]	<b>253.60</b> [60.53]	<b>-128.09</b> [77.68]
Transport cost		<b>-0.076</b> [0.022]	-0.003 [0.019]	<b>-0.139</b> [0.025]	0.031 [0.022]	<b>-0.186</b> [0.032]		-0.030 [0.021]	<b>0.032</b> [0.016]	-0.031 [0.027]	-0.010 [0.025]	<b>-0.199</b> [0.039]
Adj. R2	0.99	0.99	0.99	0.99	0.99	0.99	0.99	0.99	0.99	0.99	0.99	0.99
SER	0.048	0.038	0.031	0.043	0.033	0.055	0.037	0.033	0.027	0.042	0.039	0.061
N	138	138	138	138	138	138	138	138	138	138	138	138

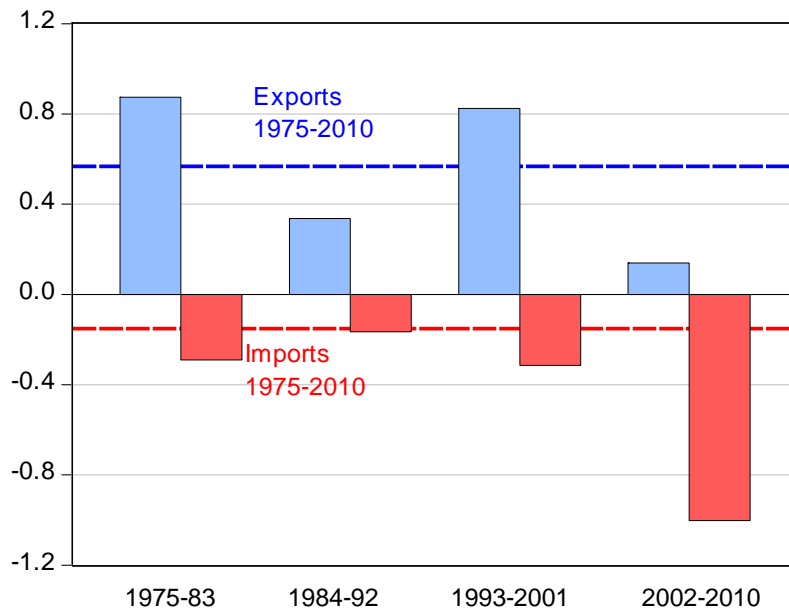
**Notes:** Point estimates and HAC standard errors for OLS and DOLS in [brackets].. SER is standard error of regression. N is number of observations. Regressions include 2 leads and 4 lags of the first difference terms of the right hand side variables. **[bold face]** indicates significance at the 10% level.



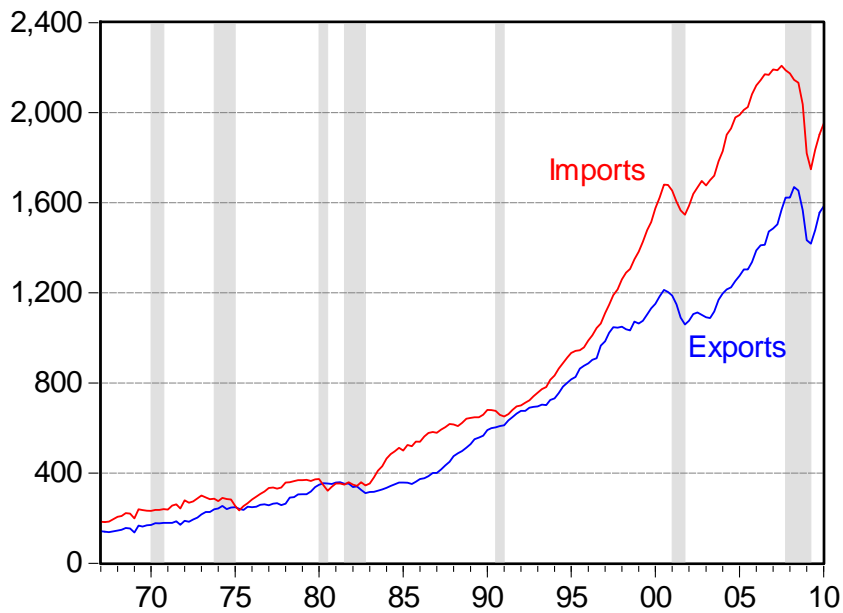
**Figure 1:** Net Exports of goods and services to GDP ratio, SAAR (left axis) and log US dollar real exchange rate against broad basket of currencies (right axis). Shaded areas denote recession dates, assumes last recession ended at 2009Q2. Source: BEA (2010Q1 2<sup>nd</sup> release) and NBER.



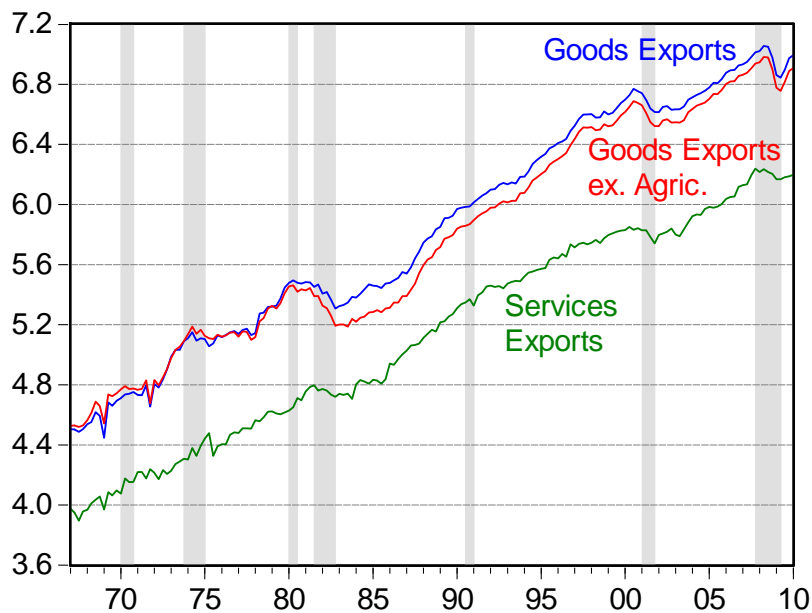
**Figure 2:** Income Export (EXPY) and Import (IMPY) Elasticities for Subperiods. Source: Columns [2] and [6] from Table 1, and DOLS regressions on the indicated subperiod.



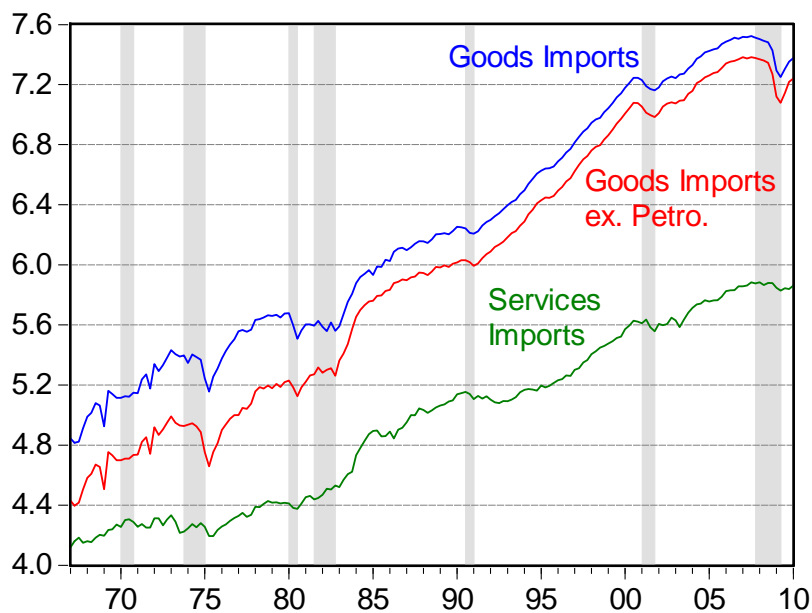
**Figure 3:** Price Export (EXPY) and Import (IMPY) Elasticities for Subperiods. Source: Columns [2] and [6] from Table 1, and DOLS regressions on the indicated subperiod.



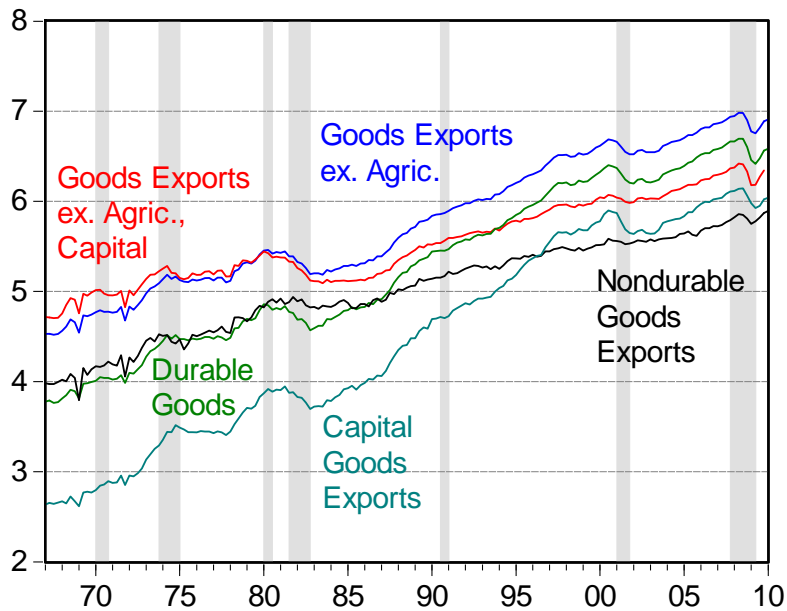
**Figure 4:** Real Exports and Imports of Goods and Services, in billions 2005 Ch\$ (SAAR). Gray shading denotes NBER recession dates. Dashed lines denote beginning and end of regression sample. Source: BEA (2010Q1 2nd release).



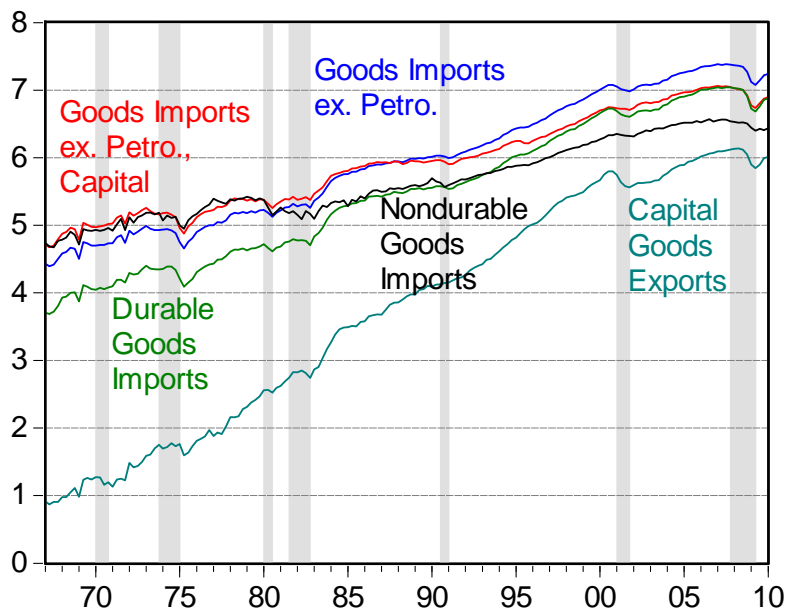
**Figure 5:** Log Real Exports of Goods, Goods ex Agricultural Goods, Goods ex Agricultural and Capital Goods, billions of 2005 Ch.\$, SAAR. Source: BEA (2009Q3 2nd release), and author's calculations.



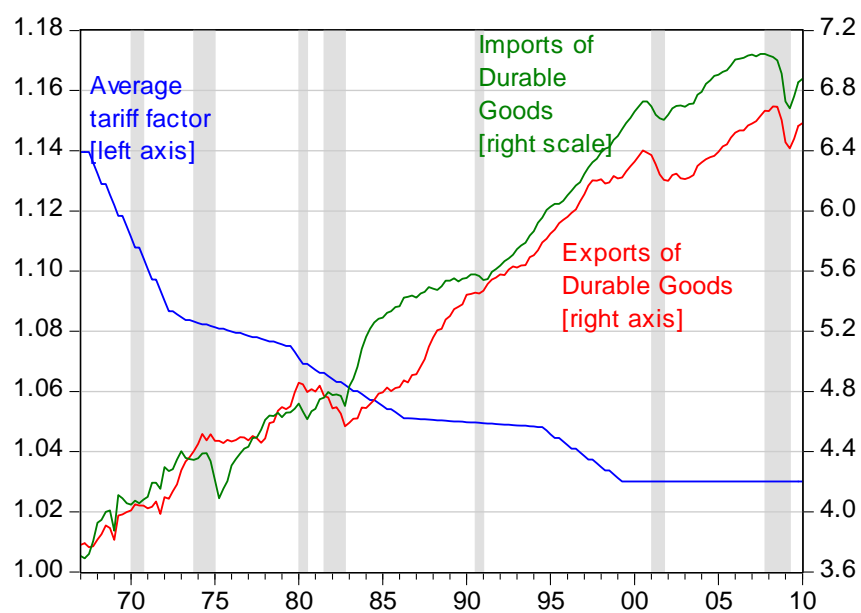
**Figure 6:** Log Real Imports of Goods, Goods ex Petroleum, and Services, billions of 2005 Ch.\$, SAAR. Source: BEA (2010Q1 2nd release), and author's calculations.



**Figure 7:** Real Exports of Goods ex. Agriculture, ex. Agriculture and Capital, of Durable Goods, of Nondurable Goods, and of Capital Goods, billions of 2005 Ch.\$, SAAR.  
Source: BEA (2010Q1 2nd release).



**Figure 8:** Real Imports of Goods ex. Petroleum, ex. Petroleum and Capital, of Durable Goods, of Nondurable Goods, and of Capital Goods, billions of 2005 Ch.\$, SAAR.  
Source: BEA (2010Q1 2nd release).



**Figure 9:** Average Tariff Factor, and Exports of Durable Goods and Imports of Durable Goods, in 2005 Ch.\$.. Source: Kei-Mu Yi, BEA and author's calculations.

## Appendix 1: Data Sources and Description

### *Exchange Rate Indices*

- US “broad” trade weighted exchange rate (CPI deflated). Source: Federal Reserve Board website, [http://www.federalreserve.gov/releases/H10/Summary/indexbc\\_m.txt](http://www.federalreserve.gov/releases/H10/Summary/indexbc_m.txt) . Data accessed June 2010. See Loretan (2005) for details.

### *Trade Flows, Economic Activity*

- Real and nominal imports and exports of goods and services, and gross domestic product (2005 chain weighted dollars). Source: BEA, 2010Q1 2nd release.
- Goods exports ex Agricultural and Capital Goods, and Goods imports ex Oil and Capital Goods calculated using Tornqvist approximation. See Whelan (2000) for an explanation of the procedure.
- For foreign demand: Rest-of-World GDP (base year). U.S. exports weighted rest-of-world GDP, ends 2009Q4. Source: personal communication, Federal Reserve Board, series obtained May 2010.
- For foreign supply: Rest-of-World GDP. U.S. import weighted growth rate for Rest-of-World, using annual data from *International Financial Statistics*, *World Economic Outlook*, *Asian Development Bank*, and *World Development Indicators*. Weights are listed at <http://www.federalreserve.gov/releases/h10/Weights/> .
- For home supply. For US, manufacturing industrial production, seasonally adjusted. Source: BEA via St. Louis Fed, accessed June 2010.
- For home supply: U.S. capital stock. US net private nonresidential capital stock from BEA. 2009 data extrapolated using a regression of net investment on gross investment and lagged capital stock (in logs, over 2001-08 period).
- Transport costs. Proxied by log oil price deflated by core US CPI. Source: BLS and Fed via St. Louis Fed, accessed June 2010.

### *Tariffs*

- Tariff rates, average of U.S., Japan and European Union, provided by Kei-Mu Yi, and described in Yi (2003). Annual data interpolated by moving average to create quarterly data.